

Weak-form Efficiency and Causality Tests in Chinese Stock Markets

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China has two major stock exchanges, the Shanghai and the Shenzhen exchanges. Each of these exchanges trades two types of shares, type “A” and type “B” shares. Type “A” shares are available to domestic investors only and type “B” shares are available to foreign investors. This article tests for the weak-form efficiency in these markets and explores the statistical relationships and causality among these Chinese stock markets with each other and with the U.S. and Hong Kong stock markets. The results indicate the existence of (1) a weak-form efficiency in the market for “A” shares but not “B” shares, (2) statistically weak linkages between the Chinese markets, (3) a weak causal effect from the Hong Kong to the four Chinese markets, and (4) a strong causal effect from U.S. stock market to all four Chinese stock markets and the Hong Kong Stock market, particularly during the second period of the sample. These results support the assertion that the Chinese stock markets are becoming more integrated to the global economy (JEL G15).

Keywords: Chinese stock markets, Granger causality tests, Hong Kong stock market, market efficiency.

I. Introduction

The Shanghai and Shenzhen are the two stock exchanges in the People’s Republic of China. Shanghai started in 1990 and Shenzhen in 1991. The exchanges operate on weekdays 9:30 a.m. to 11:00 a.m. and 1:30 p.m.

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to 3:00 p.m.¹ Each exchange trades two types of shares, known as type “A” and type “B” shares. Type “A” shares are denominated in Chinese yuan. They are sold to Chinese nationals only. Type “B” shares traded in the Shanghai exchange are denominated in U.S. dollars and those traded in the Shenzhen exchange are denominated in Hong Kong dollars. Type “B” shares are sold to investors from Hong Kong, Taiwan, Macau, and other foreign countries.² Both the Shanghai and Shenzhen exchanges use computerized trading systems.

The issuance of “B” shares cannot exceed the limit set by the government, which is 25% of the total shares of a firm. According to Huang and Liu (1995), “B” shares are intended to attract foreign investments and hard currencies. “B” shares are also developed to limit the investment activities of foreigners so that Chinese stock markets will not be overly influenced by foreign speculators.

By early 1993, the Chinese stock market capitalization had risen to 20 billion U.S. dollars and continued rising to over 100 billion U.S. dollars by 1996. In May 1998, about 789 firms were listed and traded on the Shanghai and Shenzhen exchanges. The Chinese government, through the China Securities Regulatory Commission (CSRC), exerts tight control over listings and trading of shares in these exchanges. Although government regulation provides some assurance against fraud and manipulation, it discourages market forces because share listings and sales are by governmental consent, not negotiated with investment banking firms. Interestingly, only “B” share issuers must adhere to international accounting reporting standards.³

Research on the Chinese stock markets is very limited. Only a few articles investigate issues related to the Chinese stock markets. Bailey (1994), using a small sample of firms from the Shanghai and Shenzhen exchanges, investigates correlations of weekly returns of “B” shares with returns of international stock indices. He also examines the discount of “B” shares relative to “A” shares. Bailey's data cover the

1. The time difference between New York and Shanghai and Shenzhen is 12 hours during summer and 13 hours during winter.

2. Chinese firms can also issue “H” shares for listing on the Hong Kong stock exchange and “N” shares for listing on the NYSE; e.g., Huang and Liu (1995).

3. A large drop in liquidity of about 25 billion dollars occurred in mid-1994. This was the result of governmental regulations, introduced in 1993, forbidding banks from providing liquidity to market participants. Since then, Chinese authorities reversed these regulations; for additional details, see Spencer (1995), pp. 28-31.

relatively short period from mid-1992 to the beginning of 1993. He finds at least partial segmentation of “B” share markets and unexplainable discounts on “B” shares relative to “A” shares. Su and Fleisher (1996) analyze the Chinese stock markets using daily data. They report that the daily and weekly returns of both “A” and “B” shares are non-random. Moreover, they find evidence of cointegration between the “A”-share markets but not the “B”-share markets.

The October 1987 international stock market crash and subsequent mini-crashes made it clear that none of the world’s markets operates independently. The linkage and causal relationships between international markets are becoming very important to both the private and public sectors in every country. Many U.S. and other international mutual funds and pension funds include Chinese stocks in their portfolios. Knowledge of the linkages between Chinese markets and other foreign markets enables portfolio managers to make more intelligent decisions. Moreover, it helps government officials to maintain stable domestic equity markets and to avoid the potential for international market crises.

This article examines the share-price behavior, weak-form market efficiency, and causality of the Chinese stock markets with each other and with the Hong Kong and U.S. stock markets. The article employs various statistical techniques, such as random walk, and Granger causality tests to investigate the above issues.

The article is organized as follows: Section II discusses the data and presents some preliminary evidence. Section III presents the Granger causality results. Section IV presents the summary and conclusions.

II. Data and Preliminary Results

The data include 1,000 daily observations for four Chinese stock market indices and one U.S. and one Hong Kong stock index. The indices used are the Shanghai “A” (SHA), the Shanghai “B” (SHB), the Shenzhen “A” (SZA), the Shenzhen “B” (SZB), Hong Kong stock exchange index (HK), and the Dow Jones industrial average (DJ) for the U.S. All of these indices are based on closing prices. The data are obtained from Dow Jones News Retrieval Services and Datastream International and cover the period March 8, 1993, to October 31, 1996. Stock index returns are calculated using the continuously compounded formula

TABLE 1. Unit Root Tests for the Log-Values and Log>Returns of Stock Indices

Index	Augmented Dickey Fuller (ADF) Statistics	
	Levels $\ln(P_t)$	First Differences $R_t = \ln(P_t) - \ln(P_{t-1})$
SHA	-2.99**	-10.07*
SHB	-1.93	-9.01*
SZA	-.88	-8.88*
SZB	-2.12	-8.29*
HK	1.01	-10.20*
US	-1.59	-9.48*

Note: SHA is Shanghai "A", SHB is Shanghai "B", SZA is Shenzhen "A", SZB is Shenzhen "B", HK is Hong Kong Stock market, and US is for the U.S. stock market. P_t represent the value of an index at time t , \ln is the natural logarithm, and R_t represent the log return series. The ADF statistics test the hypothesis of a unit root in each series. Statistically insignificant statistics imply acceptance of the null hypothesis. *Significant at the 1% level. **Significant at the 5% level.

$$R_t = \ln\left(\frac{P_t}{P_{t-1}}\right), \quad (1)$$

where P_t is the price of the index at time t and \ln is natural logarithm. Note that $R_t = \ln(P_t - P_{t-1}) = \Delta \ln P_t$.

To identify possible time structural changes in the behavior of the data, the sample is divided into two sub-samples covering the periods March 8, 1993, to December 31, 1994, and January 1, 1995 to October 31, 1996. Various statistical tests are performed on the two sub-samples as well as the full sample. These include serial correlation tests on daily returns and weekly returns to detect whether these markets exhibit persistent relations over time. Non-normality and runs tests are also used to examine whether the serial correlation tests are influenced by non-normality of the returns.

Before estimating the linkages between different markets, a unit root test is performed for each of the six stock indices using the augmented Dickey-Fuller (ADF) methods; see Dickey and Fuller (1979, 1981). The Dickey-Fuller method involves estimating the model

$$\Delta Y_t = \alpha_0 + \alpha_1 t + \alpha_2 Y_{t-1} + \sum_{s=1}^k c_s \Delta Y_{t-s} + u_t \quad (2)$$

TABLE 2. Phillips and Ouliaris Residual-based Bivariate Cointegration Tests of the Stock Indices

Indices	Period 1 Z_t	Period 2 Z_t	Periods 1&2 Z_t
SHA-SHB	-1.92	-1.10	-2.28
SHA-SZA	-3.43*	-2.40	-3.55*
SHA-SZB	-1.93	-2.52	-2.62
SHA-US	-2.48	-1.63	-2.58
SHB-SZA	-2.01	-2.73	-.95
SHB-SZB	-3.23	-2.19	-1.98
SHB-US	-2.35	-2.70	-2.77**
SZA-SZB	-1.36	-1.83	-.79
SZA-US	-2.09	-.15	-.66
SZB-US	-1.91	-1.38	-1.47

Note: SHA is Shanghai “A”, SHB is Shanghai “B”, SZA is Shenzhen “A”, SZB is Shenzhen “B”, and US is for the U.S. stock market. The critical values for the Z_t statistics are -2.76 at the 5% level and -3.39 at the 1% level of significance. *Significant at the 1% level. **Significant at the 5% level.

where $Y_t = \ln(P_t)$ is the natural logarithm of the value of each stock index at time t , Δ is the first difference operator, and t is for time. Table 1 presents the ADF statistics for testing the null hypothesis of a unit root $H_0: \alpha_2 = 0$ against the alternative hypothesis of a stationary time-series $H_1: \alpha_2 < 0$. All ADF statistics provide support to the null hypothesis of a unit root at the one-percent level of significance. The table also presents the ADF test statistics for the log returns series, R_t . Note that, in this case, the variable $Y_t = R_t$. All ADF statistics reject the null hypothesis of a unit root for the log-return series.

Table 2 provides bivariate cointegration tests for each pair of stock indices using the Phillips and Ouliaris (1990) method.⁴ The Philips-Ouliaris method involves estimating via OLS the regression model

$$Y_t = \beta_0 + \beta_1 X_t + e_t, \quad (3)$$

where Y_t and X_t are the log values of any two stock indices and e_t is the regression residual. The series of residuals e_t for a cointegrated pair Y_t and X_t is expected to be stationary. Table 2 presents Philips-Ouliaris Z_t statistics for testing the null hypothesis of cointegration (i.e., that the

4. A pair of economic time series is cointegrated if there is a stationary long-run relationship in this pair, although in the short-run both series deviate from each other.

TABLE 3. Summary Statistics for Shanghai and Shenzhen Stock Exchanges

Statistics	Period 1	Period 2	Periods 1 & 2
SHA Market			
Annualized Daily Return	24.16%	-28.87%	-5.42%
Standard Deviation	2.82%	4.23%	3.61%
Skewness	1.8235	1.4975	1.6160
Kurtosis	25.5473	9.6699	14.2281
Max. Return	27.85%	30.85%	30.85%
Min. Return	-18.43%	-14.60%	-18.43%
Observations	447	463	910
SHB Market			
Annualized Daily Return	-16.11%	-27.97%	-18.79%
Standard Deviation	1.04%	1.93%	1.56%
Skewness	1.2159	.2619	.4201
Kurtosis	5.3796	3.8316	5.9453
Max. Return	5.74%	9.53%	9.53%
Min. Return	-3.90%	-9.54%	-9.54%
Observations	447	463	910
SZA Market			
Annualized Daily Return	3.2989	-.5198	1.3823
Standard Deviation	.2295	.0382	.1629
Skewness	1.2355	1.4892	1.3958
Kurtosis	20.1756	10.1643	13.3628
Max. Return	.25968	.295	.295
Min. Return	-.19622	-.14761	-.1962
Observations	439	456	895
SZB Market			
Annualized Daily Return	.0273	-.421	-.1539
Standard Deviation	.0387	.0126	.0286
Skewness	-.5316	-.5316	-.9353
Kurtosis	6.1211	6.1211	29.098
Max. Return	.1222	.0449	.1222
Min. Return	-.0719	-.0719	-.1668
Observations	437	449	886

Note: SHA is Shanghai "A", SHB is Shanghai "B", SZA is Shenzhen "A", and SZB is Shenzhen "B". Period 1 is from 3/8/93 to 12/31/94 and Period 2 is from 1/1/95 to 10/31/96. Data are collected from the Dow Jones News Retrieval Services and Datastream.

residuals are stationary). The Z_t statistics reject the hypothesis of cointegration for almost all pairs of stock indices. These results imply that models for testing linkage or causality in the return series do not require the inclusion of an error correction term.

Table 3 presents some summary statistics on the return series of the

TABLE 4. Serial Correlations and Ljung-Box Statistics of Daily Returns for the Shanghai and Shenzhen Stock Exchanges

	SHA	SHB	SZA	SZB
A. Serial Correlations				
Lags				
1	-.0258	.2304	-.0181	.0704
2	.0856	.1024	.0648	.0670
3	-.0086	.0168	-.0094	.0463
4	.0163	.0815	.0606	.0968
5	.0546	.0129	.0237	.0009
6	-.1106	-.0169	-.0739	-.0045
7	.0286	.0324	.0199	-.0409
8	-.0111	.0429	-.0050	.1114
9	.0720	.0205	.0587	.0271
10	-.0771	-.0035	-.0445	.0831
11	-.0124	.0347	.0104	-.0573
12	.0201	.0607	.0282	-.0152
13	-.0308	.0606	.0063	.0583
14	.0297	.0882	.0029	.0060
15	-.0764	-.0129	-.0552	.0103
16	.0525	-.0062	.0124	.0336
17	.0003	.0057	-.0079	.0396
18	.0050	.0095	-.0381	.0390
19	-.0217	-.0409	.0164	.1010
20	-.0737	.0074	-.1176	-.0031
21	.0177	.0028	.0505	.0634
22	.0048	-.0070	-.0133	.0147
23	-.0202	.0072	.0113	.0460
24	-.0037	.0234	.0321	-.0104
B. Ljung-Box Statistics				
Period 1	56.22*	48.16*	36.78**	116.31*
Period 2	22.34	66.46*	21.37	32.42
Periods 1 & 2	49.18*	85.62*	40.59**	65.27*

Note: SHA is Shanghai "A", SHB is Shanghai "B", SZA is Shenzhen "A", and SZB is Shenzhen "B". Period 1 is from 3/8/93 to 12/31/94 and Period 2 is from 1/1/95 to 10/31/96. The autocorrelation coefficients are for the entire period. Ljung-Box statistics are computed using 24 lags. *Significant at the 1% level. **Significant at the 5% level.

four Chinese stock market indices. These preliminary results indicate that, in most cases, the annualized mean daily returns are negative.⁵ This may be attributed to investor uncertainty about future firm

5. Daily returns are annualized using the formula $AR_t = 250 \times R_t$, which assumes 250 trading days per year.

TABLE 5. Serial Correlations and Ljung-Box Statistics of Weekly Returns for the Shanghai and Shenzhen Stock Exchanges

	SHA	SHB	SZA	SZB
A. Serial Correlations				
Lags				
1	.0231	.1525	.0501	-.0994
2	-.0710	.0627	-.0188	.0948
3	.0344	.0897	-.0283	.0905
4	-.1035	-.0122	-.1106	.0993
5	-.0566	.0195	.1185	.0992
6	.0438	.0658	.0537	.0992
7	-.0721	.0883	-.0396	.1665
8	.0181	-.0465	.0448	-.1595
9	-.0911	.0033	-.1958	.1195
10	-.0290	-.1613	.0189	.0299
11	.0686	-.1054	.1234	.0455
12	-.1091	-.1314	-.0334	-.0077
13	.0313	.1552	.1460	-.0081
14	.1216	.0481	.0805	.0023
15	.0145	-.0798	.0034	-.0754
16	.0136	-.1008	-.0226	-.0511
17	-.0690	-.1547	-.0528	-.0671
18	-.0078	-.0744	.0651	-.0683
19	-.1103	-.0401	-.0277	-.0541
20	-.0486	.0235	-.0950	-.0407
21	.0123	-.0293	-.0231	-.0996
22	-.0346	-.0461	-.0482	-.1095
23	-.0004	-.1363	.0477	-.0933
24	.0325	-.1803	.1265	-.0381
B. Ljung-Box Statistics				
Period 1	17.87	46.01*	30.82	34.87**
Period 2	16.41	30.77	19.64	114.12*
Periods 1 & 2	16.61	15.27	27.98	20.03

Note: SHA is Shanghai "A", SHB is Shanghai "B", SZA is Shenzhen "A", and SZB is Shenzhen "B". Period 1 is from 3/8/93 to 12/31/94 and Period 2 is from 1/1/95 to 10/31/96. The autocorrelation coefficients are for the entire period. Ljung-Box statistics are computed using 24 lags. *Significant at the 1% level. **Significant at the 5% level.

performance and lack of market liquidity stemming from governmental restrictions on banks' participation in securities lending. The standard deviations of annualized daily returns, relative to those for the U.S., are moderate in size.⁶ Unlike the other three markets, stock returns on

6. The average standard deviation of annual returns for firms traded on the OTC during

Shenzen “B” are negatively skewed and exhibit high kurtosis values.

Tables 4 and 5 report the serial correlations and Ljung and Box (1978) test statistics for 24 lags for daily and weekly returns, respectively. The Ljung-Box statistics are mostly significant at the 5% level, indicating the presence of significant serial correlation in the daily return series in all four markets. Autocorrelation in securities markets does not necessarily imply market inefficiency; see Leroy (1973) and Levich (1979). However, the relatively large reported values for autocorrelation may indicate a violation of the weak form of market efficiency, since investors may be able to exploit serial correlation for profit making. Except for Shanghai “B”, the magnitude of serial correlation in the remaining three markets decreases during the second period, implying that the Chinese markets are moving closer to becoming efficient. As expected, the degree of serial correlations is weaker for weekly returns; see table 5. Most Ljung-box test statistics are statistically insignificant at the 1%. Note that because of transaction cost, low autocorrelation values are harder to exploit for profit making.

Another factor possibly affecting serial correlation is non-normality of the daily returns distribution. Table 6 summarizes runs tests for price changes between 1 to 5 trading days for the full sample and the two sub-samples. Of the 60 runs tests (20 per sample period), only 5 are non-significant in the overall sample period while 7 are non-significant in the early period and 15 in the latter period. This also suggests increasing market efficiency over time. Five of the six one-day returns calculated for the markets of “A” shares have greater actual than expected runs. This implies a negative serial correlation supporting some of the findings in Table 4. However, of a total of more than 30 runs tests on the markets of “A” shares, only 6 have greater actual than expected runs, but only 2 of those are statistically significant. More importantly, all 30 runs for the markets of “B” shares show fewer actual than expected runs. Nineteen of these runs are statistically significant. These results are in line with the positive serial correlations of “B” shares in table 4.

III. Granger Causality Tests

Granger's (1969) causality tests are used to investigate lead-lag relationships among the four Chinese markets and with certain international markets such as those of the U.S. and Hong Kong. Let Y_t

the period 1960-84 was 22.4%; see Ibbotson and Brinson (1987), p. 83.

TABLE 6. Run Tests on Returns for the Shanghai and Shenzhen Stock Exchanges

Markets	Period 1				Period 2				Periods 1 & 2			
	AR	ER	SD	Z	AR	ER	SD	Z	AR	ER	SD	Z
One-day return (differencing interval 1)												
SHA	254	232.26	10.74	2.03**	234	222.42	10.48	1.10	490	454.90	15.05	2.33**
SHB	197	232.50	10.75	-3.30*	203	222.42	10.48	-1.85	396	454.97	15.05	-3.92*
SZA	249	230.31	10.65	1.76	216	221.63	10.48	-.54	457	451.14	14.95	.39
SZB	212	232.19	10.73	-1.88	198	214.09	10.16	-1.58	398	451.66	14.99	-3.58*
Two-day return (differencing interval 2)												
SHA	103	116.45	7.58	-1.77	101	110.30	7.34	-1.27	203	226.81	10.60	-2.25**
SHB	96	116.32	7.57	-2.68*	95	109.60	7.29	-2.00**	191	225.99	10.56	-3.31*
SZA	118	113.85	7.41	.56	93	109.84	7.34	-2.30**	195	224.64	10.52	-2.82*
SZB	87	116.39	7.58	-3.88*	88	105.63	7.08	-2.49**	173	223.63	10.50	-4.82*
C. Three-day return (differencing interval 3)												
SHA	72	76.95	6.12	-.81	71	75.50	6.08	-.74	138	152.30	8.68	-1.65
SHB	70	77.24	6.14	-1.18	63	72.28	5.82	-1.59	135	150.48	8.57	-1.81
SZA	65	75.12	5.97	-1.69	60	73.52	5.96	-2.27**	132	149.47	8.54	-2.04**
SZB	61	77.24	6.14	-2.64*	64	73.41	5.99	-1.57	116	150.48	8.63	-4.00*

(Continued)

TABLE 6. (Continued)

D. Four-day return (differencing interval 4)												
SHA	5	57.24	5.22	-.43	53	56.28	5.22	-.63	101	113.70	7.46	-1.70
SHB	46	57.77	5.27	-2.23**	53	54.87	5.09	-.37	99	112.38	7.38	-1.81
SZA	58	57.97	5.29	.00	48	53.84	5.04	-1.16	108	113.12	7.46	-.69
SZB	40	58.46	5.33	-3.46*	46	52.81	4.98	-1.37	80	111.51	7.38	-4.27*
E. Five-day return (differencing interval 5)												
SHA	42	46.05	4.70	-.86	40	44.22	4.61	-.92	80	89.87	6.62	-1.49
SHB	38	45.84	4.67	-1.68	37	41.97	4.36	-1.14	79	88.46	6.52	-1.45
SZA	47	46.23	4.71	.16	36	42.43	4.41	-1.46	86	89.27	6.58	-.50
SZB	27	46.45	4.74	-4.11*	32	43.49	4.58	-2.51**	58	89.47	6.63	-4.75*

Note: SHA is Shanghai “A”, SHB is Shanghai “B”, SZA is Shenzhen “A”, and SZB is Shenzhen “B”. AR is the total number of runs, ER is the expected runs, SD is the standard deviation of runs, $Z = (\text{Return} - \text{ER})/\text{SD}$. *Statistically significant at the 1% level. **Statistically significant at the 5% level.

and X_t be the return series for two different markets. The following standard definition for Granger's causality is used:

Returns in market "X" cause returns in market "Y" if and only if the past values of X_t combined with past values of X_t improve the prediction of Y_t .

The empirical procedure for testing Granger's causality from market X to markets Y involves the following steps. Estimate an autoregressive (AR) model of order p

$$Y_t = \alpha_0 + \alpha_1 Y_{t-1} + \dots + \alpha_p Y_{t-p} + e_t, \quad (4)$$

where Y_t represents the return in market "Y" at time t . The lag-length p of (4) is determined using Akaike's minimum final prediction error (FPE). According to the FPE criterion, the optimal value of p is the one that minimizes the FPE statistic; see Hsiao (1981) for the details.

Next, estimate an expanded version of equation 4 which includes past values of the X_t series. That is, estimate

$$Y_t = \alpha_0 + \alpha_1 Y_{t-1} + \dots + \alpha_k Y_{t-p} + \beta_1 X_{t-1} + \dots + \beta_k X_{t-q} + u_t. \quad (5)$$

The lag-length of X_t is also determined using the FPE criterion, given the value of p from the estimation of equation 4.

The F-statistic below

$$F = \frac{(SSR_1 - SSR_2)/q}{SSR_2/(T - p - q - 1)}, \quad (6)$$

where SSR_1 is the sum of squared residuals of (4), SSR_2 is the sum of squared residuals of (5), T is the sample size, and p and q are as defined previously, provides a formal test for causality. Specifically, a statistically significant value indicates causality from market "X" to market "Y".

Table 7 presents the optimal values for p and q and F-values for testing bivariate causality between all possible pairs of the Chinese stock markets. It also presents similar statistics for possible bivariate

causality from the U.S. and Hong Kong stock markets to all four Chinese markets.

Panel A presents the statistics for bivariate causality among the four Chinese markets. The results show that Shanghai “A” has no apparent influence on Shenzhen “A”. All F-values are statistically insignificant. A causal relation exists from Shanghai “A” to Shanghai “B” in the second period and from Shanghai “A” to Shenzhen “B” in the first period. Shenzhen “A” appears to lead Shanghai “A” during the first period, but not in the second period, suggesting a fading leadership over time. This may be attributed to the late paramount leader Deng's efforts to establish Shenzhen as the “pilot” or model center of economic development and reform for China. Shanghai “B” influences the other three markets during both periods. Shenzhen “B” influences Shanghai “A” and Shenzhen “A” during the first period and Shanghai “B” during the second period, only. In sum, these results suggest a causal relationship from Shanghai “B” to all Chinese markets and from Shanghai “A” and Shenzhen “B” back to Shanghai “B”. The causal relationships from the “B” markets to the “A” imply that foreign markets exert a significant influence on the markets open only to Chinese nationals.⁷

Panel B of table 7 presents the statistics for testing causality from the Hong Kong stock market to the four Chinese markets. The results indicate a strong causal relation from the HK market to Shanghai “B” during the first period and to Shanghai “A” during the second period, only. No causal relation exists between the HK market with Shenzhen “A” or Shenzhen “B”. The latter result may be attributed to the proximity of Shenzhen to Hong Kong, which may have made it a *de facto* Chinese market.

Panel C of table 7 demonstrates that the U.S. stock market exhibit a strong causal relation to all four Chinese markets as well as the Hong Kong market during the second period. These relations are, however, weak for Shanghai “B”, Shenzhen “A”, and Shenzhen “B” during the first period. This suggests that the four Chinese markets are gradually being integrated into global financial markets.

The following explanation could be offered regarding the gradual integration of Chinese markets into the world markets. There is an increasing interest of international investors towards the Chinese markets. For example, there were no U.S. mutual funds investing in

7. It is important to note that shares are not cross-listed in the four Chinese markets.

TABLE 7. Bivariate Causality Tests Among the Four Chinese Stock Markets and from the Hong Kong and the U.S. Stock Markets to the Chinese Stock Markets

	Period 1			Period 2			Periods 1 & 2		
	<i>p</i>	<i>q</i>	F-value	<i>p</i>	<i>q</i>	F-value	<i>p</i>	<i>q</i>	F-value
A. Causality Among the Chinese Markets									
SHA to SHB	8	1	.48	7	9	2.49**	8	6	1.81
SHA to SZA	7	1	.54	3	1	.06	6	1	.32
SHA to SZB	8	4	4.30*	3	1	1.92	2	7	3.87*
SHB to SHA	9	2	2.11**	3	4	2.13**	9	2	3.73*
SHB to SZA	7	1	3.11*	3	4	2.01**	6	2	3.57*
SHB to SZB	8	7	9.71*	3	8	2.50**	2	8	4.66*
SZA to SHA	9	2	9.45**	3	1	.30	9	5	4.57**
SZA to SHB	8	1	2.98**	7	1	.18	8	5	2.36**
SZA to SZB	8	3	7.82**	3	1	3.94	2	7	4.36**
SZB to SHA	9	1	3.76*	3	2	1.27	9	1	.02
SZB to SHB	8	1	1.81	7	2	7.60*	8	2	4.63*
SZB to SZA	7	1	2.96*	3	1	1.02	6	1	.01

(Continued)

TABLE 7. (Continued)

	Period 1			Period 2			Periods 1 & 2		
	<i>p</i>	<i>q</i>	F-value	<i>p</i>	<i>q</i>	F-value	<i>p</i>	<i>q</i>	F-value
B. Causality from HK to the Chinese Markets									
HK to SHA	9	3	1.40	3	6	2.10**	9	6	2.46**
HK to SHB	8	5	6.39*	7	1	.02	8	5	7.49*
HK to SZA	7	1	.05	3	1	.07	6	1	.04
HK to SZB	8	4	1.56	3	1	1.34	2	1	.34
C. Causality from the US to the Chinese markets									
US to SHA	9	1	2.13**	3	2	2.71**	9	4	1.52
US to SHB	8	1	1.02	7	1	2.99**	8	1	3.60*
US to SZA	7	2	1.50	3	1	2.55**	6	1	.44
US to SZB	8	1	.20	3	1	2.23**	2	1	1.46
US to HKS	7	2	31.19*	8	1	25.0*	7	3	53.83*

Note: SHA is Shanghai “A”, SHB is Shanghai “B”, SZA is Shenzhen “A”, SZB is Shenzhen “B”, HK is Hong Kong Stock Exchange, and US is for the U.S. stock market. Period 1 is from 3/8/93 to 12/31/94 and Period 2 is from 1/1/95 to 10/31/96. *Significant at the 1% level. **Significant at the 5% level.

Chinese companies before 1993. By the end of 1997, there were many U.S. mutual funds companies that invested in China. Moreover, an increasing number of Chinese companies are now dually listed in U.S. markets. Therefore, Chinese markets get feedback of the movements from these dual listing stocks.

IV. Summary and Conclusions

This article tests for the weak-form efficiency in the four Chinese stock markets of Shanghai “A”, Shanghai “B”, Shenzhen “A”, and Shenzhen “B” and explores the presence of causality influences between these markets as well as the U.S. and Hong Kong stock markets.

The Ljung-Box test statistics indicate the presence of significant serial correlation in the daily return series in all four Chinese markets. This may be interpreted as a violation of the weak form of market efficiency, since investors may be able to exploit serial correlation for profit making. Except for Shanghai “B”, the magnitude of serial correlation in the remaining three markets decreases during the second period of the data, implying that the Chinese markets are moving closer to becoming efficient.

Granger causality tests indicate a causal relationship from Shanghai “B” to the other three Chinese stock markets and from Shanghai “A” and Shenzhen “B” back to Shanghai “B”. The causal relationships from the “B” stock markets to the “A” stock markets imply that foreign markets exert a significant influence on the markets open only to Chinese nationals. The Hong Kong stock market exerts a significant causal relationship on Shanghai “B” during the first period and on Shanghai “A” during the second period. The U.S. stock market exhibits a strong causal relation to all four Chinese markets as well as the Hong Kong market, especially during second period. These results suggests that the four Chinese markets are gradually being integrated into the global economy.

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